

ON A SKEW STABLE LÉVY PROCESS

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ABSTRACT. The skew Brownian motion is a strong Markov process which behaves like a Brownian motion until hitting zero and exhibits an asymmetry at zero. We address the following question: what is a natural counterpart of the skew Brownian motion in the situation that the noise is a stable Lévy process with finite mean and infinite variance. We define a skew stable Lévy process X as the limit of a sequence of stable Lévy processes which are perturbed at zero. We point out a formula for the resolvent of X and show that X is a solution to a stochastic differential equation with a local time. Also, we provide a representation of X in terms of Itô's excursion theory.

1. INTRODUCTION AND MAIN RESULTS

A skew Brownian motion (SBM) with parameter $p \in [0, 1]$ appears in the book [16] as a diffusion that behaves like a Brownian motion until hitting 0 and whose excursions select the positive or negative sign with probabilities p and $1 - p$, respectively. There are numerous alternative descriptions and generalizations of the SBM. Portenko constructed an SBM as a solution to a stochastic differential equation (SDE) with a generalized drift or a diffusion with a semipermeable membrane [31, 32]. Harrison and Shepp [13] proved that an SBM is a strong solution to an SDE with a local time drift, which can be thought of as Dirac's delta function drift in Portenko's framework. Harrison and Shepp also constructed an SBM as an appropriate limit of random walks perturbed at 0, see also [14, 23, 24, 28] and references therein. Walsh [35] described an SBM with the help of a martingale problem, see also [2]. The detailed review on the SBM is presented in [20].

In this paper we address the following question: what is a natural analogue of the SBM in the case of a stable noise? For $\alpha \in (1, 2)$, denote by $U_\alpha := (U_\alpha(t))_{t \geq 0}$ a symmetric α -stable Lévy process with characteristic function

$$\mathbb{E} \exp(izU_\alpha(t)) = \exp(-t|z|^\alpha), \quad z \in \mathbb{R}, t \geq 0.$$

This process hits any point with probability 1, see, for instance, [3, p. 63] or [34, Example 43.42]. We are going to construct a Feller process that behaves like U_α until hitting 0 and have some "asymmetry" at the origin. To understand what could be a natural asymmetry at 0 we attempt to find analogies with the SBM construction. Note that U_α does not hit 0 by a single jump and cross the zero level infinitely often before hitting 0, see [36, Theorem 6.4]. It does not also exit 0 by a jump and changes sign infinitely often on exiting 0 [34, Theorem 47.1]. Thus, unlike in the case of SBM, selecting signs of the excursions becomes an issue.

The transition probability density function of the SBM with parameter p at time $t > 0$ is given by

$$(x, y) \mapsto \varphi_t^{(2)}(x - y) + (2p - 1)\operatorname{sgn}(y)\varphi_t^{(2)}(|x| + |y|), \quad x, y \in \mathbb{R},$$

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where $\varphi_t^{(2)}$ is the density of a centered normal distribution with variance t . For $t > 0$, denote by $\varphi_t^{(\alpha)}$ the density of the random variable $U_\alpha(t)$, $\alpha \in (1, 2)$. It turns out that, for $p \in [0, 1]$,

$$(x, y) \mapsto \varphi_t^{(\alpha)}(x - y) + (2p - 1)\operatorname{sgn}(y)\varphi_t^{(\alpha)}(|x| + |y|), \quad x, y \in \mathbb{R}$$

is the transition density of a Markov process, whose martingale characterization is given in [30]. Now we provide an informal construction of this process and note that it does not behave like U_α outside 0. The jumps of the process are governed by the intensity function proportional to $|x|^{-(1+\alpha)}$, $x \in \mathbb{R}$. These are accumulated until the process changes sign. At the epoch of sign change a new sign is selected positive with probability p and negative with probability $1 - p$.

Existence and uniqueness of a strong solution to an SDE with a local time were proved by Harrison and Shepp with the help of Tanaka's formula and Nakao's theorem. This technique fails in the case of an α -stable noise, and so do arguments related to space- and time-change transforms, which are efficient for one-dimensional diffusions with local times, see [19]. When constructing a diffusion with a generalized drift Portenko uses a partial differential equations approach. An essential part of his proof is based on the result dealing with a jump of the normal derivative of a single layer potential. It turns out that a similar result, in which the derivative has to be replaced with some nonlocal operator, holds true for a potential generated by the process U_α . Unfortunately, such an approach leads to strongly continuous semigroups without nonnegativity condition [21, 25].

We also mention here several recent results on strong solutions of SDEs with singular drifts and additive fractional Brownian motion noise having a small Hurst parameter, see [1, 7] and references therein. These results are derived under the assumption that the noise has a local time, which is sufficiently smooth with respect to the spatial parameter. Observe that the α -stable Lévy process U_α does not enjoy such a property.

To define the asymmetry at 0 in a natural way, we shall use the approach of Harrison and Shepp. Specifically, we construct perturbations of the α -stable process U_α as follows. Let ζ_1, ζ_2, \dots be independent copies of a random variable ζ , which are independent of U_α . Assume that $\zeta \neq 0$ almost surely (a.s.). With these at hand, we define the process $X_\zeta := (X_\zeta(t))_{t \geq 0}$ which satisfies $X_\zeta(0) = x \neq 0$ and has the same increments as U_α on the time intervals where X_ζ does not "touch" 0. Upon the k th touch of 0 the process X_ζ has jump ζ_k . To make the previous discussion formal, put

$$\begin{aligned} \sigma_0 &= 0, \quad \sigma_{k+1} := \inf\{t > \sigma_k : \zeta_k + U_\alpha(t) - U_\alpha(\sigma_k) = 0\}, \\ X_\zeta(t) &:= x + U_\alpha(t), \quad t \in [0, \sigma_1), \\ X_\zeta(t) &:= \zeta_k + U_\alpha(t) - U_\alpha(\sigma_k), \quad t \in [\sigma_k, \sigma_{k+1}), \quad k \geq 1. \end{aligned} \tag{1}$$

Now we want to successively decrease the perturbations $(\zeta_k)_{k \geq 1}$ of U_α . To this end, for each positive integer n , replace $(\zeta_k)_{k \geq 1}$ with $(\zeta_k/n)_{k \geq 1}$, then define the processes $X_{\zeta/n}$ and send $n \rightarrow \infty$. Note that each particular perturbation tends to 0. Nevertheless, the smaller the jump from 0 is, the smaller the return time to 0 is. As a consequence, the number of visits to 0 increases as n grows.

In this paper we aim at finding a distributional limit (that we denote by X) for the sequence $(X_{\zeta/n})_{n \geq 1}$ and investigating its properties. In particular, we shall point out the resolvent, the entrance law, the excursions measure and an SDE that X satisfies. Related to our investigation are the papers [18, 38, 39], in which some invariance principles are obtained in terms of convergence of the excursions measures.

Before formulating our results we note that, for $t > 0$ and large n , the value $X_{\zeta/n}(t)$ is the sum of $U_\alpha(t)$ and a random number (depending on t) of small perturbations from the collection $(\zeta_k/n)_{k \geq 1}$. Intuitively, if the distribution of ζ is light-tailed, the contribution of the perturbations should be negligible. We shall show below that this (trivial) situation occurs whenever $E|\zeta| < \infty$. Assume now that the opposite situation prevails, which particularly means that $E|\zeta| = \infty$. Then, to make the sum of perturbations, properly normalized, convergent, it is natural to assume that the distribution of ζ belongs to the domain of attraction of a β -stable distribution with $\beta \in (0, 1)$ (here, β could have been equal to 1; however, we do not treat this case in the present paper). This means that the variables $\zeta_1 + \dots + \zeta_n$, properly normalized and centered, converge in distribution to a random variable with a β -stable distribution. It is known that this happens if, and only if, the function $x \mapsto P(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$ and the limits

$$c_\pm := \lim_{x \rightarrow +\infty} \frac{P(\pm\zeta > x)}{P(|\zeta| > x)} \quad (2)$$

exist and satisfy $c_- + c_+ > 0$.

Given next are some notation to be in force throughout the paper. For a process Y , denote by σ or $\sigma(Y)$ the first hitting time of 0, that is,

$$\sigma(Y) = \sigma := \inf\{t > 0 : Y(t) = 0\}.$$

For bounded measurable functions $f : \mathbb{R} \rightarrow \mathbb{R}$, put

$$V_\lambda f(x) := E^x \int_0^\sigma e^{-\lambda s} f(U_\alpha(s)) ds,$$

so that V_λ is the resolvent of U_α killed at 0. We write $D := D[0, \infty)$ for the Skorokhod space of càdlàg functions defined on $[0, \infty)$. We always assume that the space D is endowed with the J_1 -topology. For a measurable function $f : \mathbb{R} \rightarrow \mathbb{R}$ and a measure ν on \mathbb{R} we write $\langle \nu, f \rangle$ for $\int_{\mathbb{R}} f(x) \nu(dx)$ provided that the integral is well-defined. In particular, if ν is a probability measure, then $\langle \nu, f \rangle = Ef(\tau)$, where τ is a random variable with distribution ν .

We are ready to formulate our main results.

Theorem A. *Assume that either the function $x \mapsto P(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$, $\beta \in (0, 1)$ and (2) holds, or $E|\zeta| < \infty$. If $X_{\zeta/n}(0)$ converges in distribution as $n \rightarrow \infty$ to some random variable ξ , then the processes $X_{\zeta/n}$ converge in distribution on D to a Feller process X starting at ξ .*

(a) *If $\beta < \alpha - 1$, then the resolvent of X is given by*

$$R_\lambda f(x) = V_\lambda f(x) + E^x e^{-\lambda \sigma(U_\alpha)} \frac{\langle \eta^*, V_\lambda f \rangle}{\langle \eta^*, V_\lambda 1 \rangle} \quad (3)$$

for bounded measurable f , where η^ is a measure defined by*

$$\eta^*(dx) = (c_- \mathbb{I}_{(-\infty, 0)}(x) + c_+ \mathbb{I}_{(0, \infty)}(x)) |x|^{-(1+\beta)} dx, \quad x \in \mathbb{R}, \quad (4)$$

and the constants c_\pm are given in (2).

(b) *If $\beta > \alpha - 1$ or $E|\zeta| < \infty$, then $X(t) = \xi + U_\alpha(t)$, $t \geq 0$, where ξ and U_α are independent.*

One of the standing assumptions of the previous theorem is $\alpha \in (1, 2)$. Put formally $\alpha = 2$, so that the noise becomes a Brownian motion W , say. Then, under the assumption $E|\zeta| < \infty$, a counterpart of the limit process in Theorem A, still denoted by X , is a SBM which solves the SDE

$$dX(t) = dW(t) + \gamma dL(t). \quad (5)$$

Here, $\gamma = \frac{E\zeta}{E|\zeta|} \in [-1, 1]$, and L is a two-sided local time of X at 0. The claim can be justified with the help of arguments given in [8] or [14]. Alternatively, this can be shown along the lines of the proof of Theorem A.

Although a resolvent uniquely determines the corresponding Feller process, formula (3), being rather implicit, does not shed much light on the properties of X . As a remedy, we characterize a Feller process with resolvent (3) as a solution to an SDE.

Theorem B. *Assume that the function $x \mapsto P(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$, $\beta \in (0, \alpha - 1)$ and (2) holds. Let X be a Feller process with resolvent (3). Then X is a weak solution to the SDE*

$$X(t) = X(0) + U_\alpha(t) + S_\beta(L_0^X(t)), \quad t \geq 0. \quad (6)$$

Here, L_0^X is a Blumenthal-Gettoor local time of X at 0, S_β is a β -stable Lévy process which is independent of U_α and has the Lévy measure η given by

$$\eta(dx) = C(c_- \mathbb{I}_{(-\infty, 0)}(x) + c_+ \mathbb{I}_{(0, \infty)}(x)) |x|^{-(1+\beta)} dx, \quad x \in \mathbb{R}, \quad (7)$$

where the constants c_\pm are given in (2),

$$C := \left(\int_0^\infty E^x(1 - e^{-\sigma(U_\alpha)}) \eta(dx) \right)^{-1} = \frac{\beta \sin \frac{\pi(\beta+1)}{\alpha}}{(c_- + c_+) \Gamma(1 - \beta) \cos \frac{\pi\beta}{2} \sin \frac{\pi}{\alpha}}$$

and Γ is the Euler gamma function.

Furthermore, the process X has a zero sojourn at 0 with probability 1.

Remark 1.1. The definition of the Blumenthal-Gettoor local time can be found in [5, Theorem 2.3, Chapter V Section 2] or Section 3 below.

Remark 1.2. The weak solution in Theorem A is a triple (X, U_α, S_β) , with all components being defined on a common probability space, which satisfies equality (6) a.s. Here, the components are as defined in Theorem A. In particular, U_α and S_β are independent.

While not discussing a filtration, we only mention that it follows from the construction that the processes U_α and $(S_\beta(L_0^X(t)))_{t \geq 0}$ are $(\mathcal{F}_t^X)_{t \geq 0}$ -adapted, where $(\mathcal{F}_t^X)_{t \geq 0}$ is a filtration generated by X and augmented by events of probability 0. Observe that the process S_β is not $(\mathcal{F}_t^X)_{t \geq 0}$ -adapted.

Comparing equations (5) and (6) we find it reasonable to call the process X with resolvent (3) a skew α -stable Lévy process.

Equation (5) has a unique solution if $|\gamma| \leq 1$ and has no solution if $|\gamma| > 1$, see [13]. An interesting problem is to find a counterpart of the parameter γ for equations like (6). Theorems C and D given next provide a solution to the problem as well as a description of the corresponding processes with the help of Itô's excursion theory. We shall recall basic definitions and results of the theory in Section 3 below.

For $t \geq 0$, put $\eta_t^{\text{jump}} := \eta P_t^0$, where the measure η is as defined in Theorem B, and P_t^0 is the semigroup of U_α killed at 0. For $x \in \mathbb{R}$, denote by \bar{P}^x the semigroup of U_α stopped at 0. It can be checked that $\bar{P}^\eta := \int_{\mathbb{R}} P^x \eta(dx)$ is the excursion measure of a skew α -stable Lévy process. Denote by $(\eta_t^c)_{t > 0}$ and \hat{P}^{U_α} the entrance law of U_α and the corresponding excursion measure, respectively.

Theorem C. *Assume that the function $x \mapsto P(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$, $\beta \in (0, \alpha - 1)$ and (2) holds. Let $p \in [0, 1]$ and X be a Feller process having the entrance law $(p\eta_t^{\text{jump}} + (1-p)\eta_t^c)_{t > 0}$ and the corresponding excursion measure $p\bar{P}^\eta + (1-p)\hat{P}^{U_\alpha}$, where the measure η is defined in (7). Then X is a weak solution to the SDE*

$$X(t) = X(0) + U_\alpha(t) + p^{\frac{1}{\beta}} S_\beta(L_0^X(t)), \quad t \geq 0, \quad (8)$$

where L_0^X, S_β are as given in Theorem B.

Furthermore, the process X has a zero sojourn at 0 with probability 1.

Theorem D. Assume that the function $x \mapsto P(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$, $\beta \in (0, \alpha - 1)$ and (2) holds. Let $p \in [0, 1]$ and S_β be a β -stable Lévy process as defined in Theorem B. Then there exists a unique Feller process X with a zero sojourn at 0, which is a weak solution to equation (8), with U_α being an $(\mathcal{F}_t^X)_{t \geq 0}$ -adapted process.

Remark 1.3. Even though we require in Theorem D $(\mathcal{F}_t^X)_{t \geq 0}$ -adaptedness of U_α , it may follow automatically from (8) and independence of U_α and S_β .

We close the section with the list of open problems which are non-trivial even for the SBM.

- 1) Characterize a sticky skew α -stable Lévy process and the corresponding SDE in the way similar to that used for the sticky Brownian motion, see [10].
- 2) Consider a time inhomogeneous analogue of a skew stable Lévy process and describe the corresponding semigroup in terms of partial differential equations with Feller-Wenzell boundary condition at 0, see [17] for the Brownian case.
- 3) Investigate existence and uniqueness of a strong solution or a path-by-path solution in the sense of Davie [9] to equation (8). This problem is non-trivial even in the case where the noise is a Brownian motion and the process S_β is a subordinator [27], see also [29, 15].
- 4) Prove uniqueness of a weak solution to (8) among all (possibly non-Markov) solutions in the situations that the local time is defined in terms of the time spent in a neighborhood of 0, or a number of long excursions, or a number of a level crossings, etc.

The remainder of the paper is structured as follows. In Section 2 we use a resolvent technique and prove Theorem A. In Section 3 we recall some basic facts of Itô's excursion theory. In Section 4 we prove Theorems B, C and D and their generalizations.

2. CONVERGENCE OF RESOLVENTS

2.1. Discussion and limit theorem. In view of the assumption $X_\zeta(0) = x \neq 0$, the process X_ζ does not visit 0. It follows from the construction that X_ζ is a strong Markov process on $\mathbb{R} \setminus \{0\}$ with càdlàg paths. We shall investigate distributional convergence of the processes $X_{\zeta/n}$ as $n \rightarrow \infty$. Since we expect that a limit process visits 0, the machinery of Markov processes on \mathbb{R} rather than $\mathbb{R} \setminus \{0\}$ has to be exploited. To this end, we introduce an auxiliary *holding and jumping process* that spends at 0 a random period of time having an exponential distribution (exponential time, in short), then has jump ζ_k and afterwards behaves like U_α until the next visit to 0. The evolution just described then iterates, and the successive exponential times at 0 are independent and identically distributed.

Here is a formal construction. For $m > 0$, denote by τ_1, τ_2, \dots independent random variables having the exponential distribution of mean $1/m$. Assume that the sequences $(\zeta_k)_{k \geq 1}$ and $(\tau_k)_{k \geq 1}$ and the process U_α are independent. Similarly to (1), put

$$\tilde{\sigma}_0 = 0, \quad \tilde{\sigma}_{k+1} := \inf\{t > \tilde{\sigma}_k + \tau_k : \zeta_k + U_\alpha(t) - U_\alpha(\tilde{\sigma}_k + \tau_k) = 0\}, \quad k \geq 0$$

and then

$$X_{\zeta,m}(t) := 0, \quad \text{for } t \in [\tilde{\sigma}_k, \tilde{\sigma}_k + \tau_k), \quad k \geq 0$$

and

$$X_{\zeta,m}(t) := \zeta_k + U_\alpha(t) - U_\alpha(\tilde{\sigma}_k + \tau_k), \quad \text{for } t \in [\tilde{\sigma}_k + \tau_k, \tilde{\sigma}_{k+1}), \quad k \geq 0.$$

The so defined $X_{\zeta,m}$ is a Feller process on \mathbb{R} with càdlàg paths. Unlike X_ζ , the process $X_{\zeta,m}$ visits 0 and may start at 0.

Recall *Slutsky's lemma*: if $(X_n)_{n \geq 1}$ and $(Y_n)_{n \geq 1}$ are sequences of random elements in a metric space (with metric dist) which satisfy $\text{dist}(X_n, Y_n) \xrightarrow{P} 0$ as $n \rightarrow \infty$, and the elements X_n converge in distribution as $n \rightarrow \infty$ to a random element X , then the elements Y_n converge in distribution to X , too.

Note that

$$d(X_\zeta, X_{\zeta, m}) \xrightarrow{P} 0, \quad m \rightarrow \infty,$$

where d is the J_1 -metric on D and \xrightarrow{P} denotes convergence in probability. Hence, distributional convergence of $X_{\zeta/n}$ to X follows if we can show that $X_{\zeta/n, m_n}$ converges in distribution, where (m_n) is a sequence which diverges to $+\infty$ sufficiently fast as $n \rightarrow \infty$.

For $x, y \in \mathbb{R}$ and a Markov process X , denote by $P_t(x, dy)$ its transition probability function at time $t > 0$. Also, for bounded measurable functions $f : \mathbb{R} \rightarrow \mathbb{R}$, we define the semigroup

$$\mathbb{E}^x f(X(t)) := \int_{\mathbb{R}} f(y) P_t(x, dy) = \langle P_t, f \rangle(x) = P_t f(x), \quad t \geq 0, x \in \mathbb{R}$$

and the resolvent

$$R_\lambda f(x) := R_\lambda^X f(x) := \mathbb{E}^x \int_0^\infty e^{-\lambda t} f(X(t)) dt = \int_0^\infty e^{-\lambda t} P_t f(x) dt, \quad x \in \mathbb{R}.$$

Further, for $x, y \in \mathbb{R}$, denote by $P_t^0(x, dy)$ the transition probability function at time $t > 0$ for the process X killed upon the first visit to 0, that is,

$$P_t^0(x, A) = \mathbb{P}^x(X(t) \in A, t < \sigma), \quad x \in \mathbb{R}$$

for Borel sets A on \mathbb{R} . Also, we define the semigroup of the killed process

$$\mathbb{E}^x f(X(t)) \mathbb{1}_{\{t < \sigma\}} = \int_{\mathbb{R}} f(y) P_t^0(x, dy) = \langle P_t^0, f \rangle(x) = P_t^0 f(x), \quad t \geq 0, x \in \mathbb{R}$$

and its resolvent

$$V_\lambda f(x) = V_\lambda^X f(x) := \mathbb{E}^x \int_0^\sigma e^{-\lambda t} f(X(t)) dt = \int_0^\infty e^{-\lambda t} P_t^0 f(x) dt, \quad x \in \mathbb{R}.$$

In Section 1 we have used the same notation V_λ for the resolvent of the particular killed Markov process U_α . Hopefully, this does not lead to a confusion. For later use, we note that if X is a strong Markov process, then

$$\begin{aligned} R_\lambda f(x) &= \mathbb{E}^x \left(\int_0^{\sigma(X)} + \int_{\sigma(X)}^\infty \right) e^{-\lambda t} f(X(t)) dt \\ &= V_\lambda f(x) + \mathbb{E}^x e^{-\lambda \sigma(X)} \int_0^\infty e^{-\lambda t} f(X(t + \sigma(X))) dt \\ &= V_\lambda f(x) + \mathbb{E}^x e^{-\lambda \sigma(X)} R_\lambda f(0), \quad x \in \mathbb{R}. \end{aligned} \quad (9)$$

Theorem 2.1 states that the uniform convergence of resolvents entails distributional convergence of the corresponding Markov processes. We write $C_0(\mathbb{R})$ for the space of all continuous functions vanishing at $\pm\infty$, equipped with the supremum norm.

Theorem 2.1. *Let $(R_\lambda^{(n)})_{n \geq 1}$ be a sequence of resolvents of some Feller processes $(X_n)_{n \geq 1}$. Assume that, for each $f \in C_0(\mathbb{R})$ and each $\lambda > 0$,*

$$\lim_{n \rightarrow \infty} \sup_{x \in \mathbb{R}} |R_\lambda^{(n)} f(x) - R_\lambda f(x)| = 0,$$

where R_λ is the resolvent of a Feller process X . Assume also that the variables $X_n(0)$ converge in distribution as $n \rightarrow \infty$ to $X(0)$. Then

$$X_n \Rightarrow X, \quad n \rightarrow \infty$$

on D .

The proof follows from [11, Theorem 2.5, Chapter 4] and [26, Theorem 4.2, Chapter 3].

We shall write $R_\lambda^{\zeta, m}$ for the resolvent of $X_{\zeta, m}$. For bounded measurable functions $f : \mathbb{R} \rightarrow \mathbb{R}$ we intend to calculate $R_\lambda^{\zeta, m} f(0)$. Although this follows a standard pattern (see, for instance, [5, pp. 136-137]), we provide full details for completeness. Write

$$\begin{aligned} R_\lambda^{\zeta, m} f(0) &= \mathbb{E}^0 \left(\int_0^{\tau_1} + \int_{\tau_1}^{\infty} \right) e^{-\lambda t} f(X_{\zeta, m}(t)) dt = \lambda^{-1} \mathbb{E}(1 - e^{-\lambda \tau_1}) f(0) \\ &\quad + \mathbb{E} e^{-\lambda \tau_1} \mathbb{E}^\zeta \int_0^{\infty} e^{-\lambda t} f(X_{\zeta, m}(t)) dt = \frac{1}{m + \lambda} f(0) + \frac{m}{m + \lambda} \langle P_\zeta, R_\lambda^{\zeta, m} f \rangle. \end{aligned} \quad (10)$$

Using (9) we infer

$$\langle P_\zeta, R_\lambda^{\zeta, m} f \rangle = \langle P_\zeta, V_\lambda^{U_\alpha} f \rangle + \mathbb{E}^\zeta e^{-\lambda \sigma(U_\alpha)} R_\lambda f(0)$$

having utilized the equality $\mathbb{E}^\zeta e^{-\lambda \sigma(X_{\zeta, m})} = \mathbb{E}^\zeta e^{-\lambda \sigma(U_\alpha)}$. Substituting this into (10) and then solving for $R_\lambda f(0)$ yields

$$\lambda R_\lambda^{\zeta, m} f(0) = \frac{\frac{f(0)}{m} + \langle P_\zeta, V_\lambda^{U_\alpha} f \rangle}{\frac{1}{m} + \lambda^{-1} \mathbb{E}^\zeta (1 - e^{-\lambda \sigma(U_\alpha)})} = \frac{\frac{f(0)}{m} + \langle P_\zeta, V_\lambda^{U_\alpha} f \rangle}{\frac{1}{m} + \langle P_\zeta, V_\lambda^{U_\alpha} 1 \rangle}. \quad (11)$$

Theorem 2.2. *Assume that either the function $x \mapsto \mathbb{P}(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$, $\beta \in (0, 1)$ and (2) holds, or $\mathbb{E}|\zeta| < \infty$. Let $f : \mathbb{R} \rightarrow \mathbb{R}$ be any bounded measurable function.*

(a) *If $\beta < \alpha - 1$, then*

$$\lim_{n \rightarrow \infty} \lambda R_\lambda^{\zeta/n, m_n} f(0) = \frac{\langle \eta^*, V_\lambda^{U_\alpha} f \rangle}{\langle \eta^*, V_\lambda^{U_\alpha} 1 \rangle}, \quad (12)$$

where $(m_n)_{n \geq 1}$ is any sequence of positive numbers satisfying $\lim_{n \rightarrow \infty} m_n \mathbb{P}(\zeta > n) = \infty$, and η^* is the measure defined by (4).

(b) *If $\beta > \alpha - 1$ or $\mathbb{E}|\zeta| < \infty$, then*

$$\lim_{n \rightarrow \infty} R_\lambda^{\zeta/n, m_n} f(0) = R_\lambda^{U_\alpha} f(0),$$

where $(m_n)_{n \geq 1}$ is any sequence of positive numbers satisfying $\lim_{n \rightarrow \infty} m_n n^{1-\alpha} = \infty$, and $R_\lambda^{U_\alpha}$ is the resolvent of U_α .

We claim that

$$\lim_{n \rightarrow \infty} \sup_{x \in \mathbb{R}} |R_\lambda^{\zeta/n, m_n} f(x) - R_\lambda f(x)| = 0$$

for any bounded measurable $f : \mathbb{R} \rightarrow \mathbb{R}$ and any $\lambda > 0$. Here, for $x \in \mathbb{R}$,

$$R_\lambda f(x) = \begin{cases} V_\lambda^{U_\alpha} f(x) + \lambda^{-1} \mathbb{E}^x e^{-\lambda \sigma(U_\alpha)} \frac{\langle \eta^*, V_\lambda^{U_\alpha} f \rangle}{\langle \eta^*, V_\lambda^{U_\alpha} 1 \rangle}, & \text{if } \beta < \alpha - 1, \\ R_\lambda^{U_\alpha} f(x), & \text{if } \beta > \alpha - 1 \text{ or } \mathbb{E}|\zeta| < \infty. \end{cases} \quad (13)$$

To check this, observe that $V_\lambda^{U_\alpha} = V_\lambda^{X_{\zeta/n, m_n}}$, $\mathbb{E}^x e^{-\lambda\sigma(X_{\zeta, m})} = \mathbb{E}^x e^{-\lambda\sigma(U_\alpha)}$ for $x \in \mathbb{R}$. Further, $R_\lambda f(0) = \lambda^{-1} \frac{\langle \eta^*, V_\lambda^{U_\alpha} f \rangle}{\langle \eta^*, V_\lambda^{U_{\alpha-1}} \rangle}$ if $\beta < \alpha - 1$ and $R_\lambda f(0) = R_\lambda^{U_\alpha} f(0)$ if $\beta > \alpha - 1$ or $\mathbb{E}|\zeta| < \infty$. With these at hand, invoking (9) yields, for any $x \in \mathbb{R}$,

$$\begin{aligned} |R_\lambda^{\zeta/n, m_n} f(x) - R_\lambda f(x)| &= \mathbb{E}^x e^{-\lambda\sigma(U_\alpha)} |R_\lambda^{\zeta/n, m_n} f(0) - R_\lambda f(0)| \\ &\leq |R_\lambda^{\zeta/n, m_n} f(0) - R_\lambda f(0)| \rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

We have used Theorem 2.2 for the limit relation.

According to [5, Chapter V, §2, Theorem 2.8], R_λ is the resolvent of a strongly continuous probability semigroup and, as such, the resolvent of a Feller process. This observation in combination with Theorem 2.1 leads to the following.

Corollary 2.1. *Assume that the variables $X_{\zeta/n}(0)$ converge in distribution as $n \rightarrow \infty$ to a random variable ξ . Then the processes $X_{\zeta/n}$ converge in distribution on D to a Feller process X with the resolvent R_λ defined in (13) and $X(0)$ having the same distribution as ξ . In particular, if $\beta > \alpha - 1$, then the limit process is $\xi + U_\alpha$.*

Theorem A is an immediate consequence of Corollary 2.1.

2.2. Auxiliary results. In this section we prove a few preparatory results needed for the proof of Theorem 2.2. We shall treat the cases $\beta < \alpha - 1$ and $\beta > \alpha - 1$ or $\mathbb{E}|\zeta| < \infty$ in slightly different ways. As a consequence, we provide two collections of auxiliary results designed to deal with these cases. Throughout this section we write σ for $\sigma(U_\alpha)$, R_λ for $R_\lambda^{U_\alpha}$ and V_λ for $V_\lambda^{U_\alpha}$.

AUXILIARY RESULTS FOR THE CASE $\beta < \alpha - 1$. For $\lambda > 0$, the density u_λ of the resolvent kernel of U_α satisfies $u_\lambda(x, y) = u_\lambda(y - x)$, $x, y \in \mathbb{R}$, where

$$u_\lambda(x) = \frac{1}{\pi} \int_0^\infty \frac{\cos(x\theta)}{\lambda + \theta^\alpha} d\theta, \quad x \in \mathbb{R}.$$

Lemma 2.1 collects a couple of formulae to be used in what follows.

Lemma 2.1. *Let $\alpha \in (1, 2)$.*

(a) *For $\gamma \in [0, \alpha - 1)$ and $\lambda > 0$,*

$$\int_0^\infty \frac{\theta^\gamma}{\lambda + \theta^\alpha} d\theta = \frac{\Gamma(1 - \frac{\gamma+1}{\alpha}) \Gamma(\frac{\gamma+1}{\alpha})}{\alpha \lambda^{1 - \frac{\gamma+1}{\alpha}}} = \frac{\pi}{\alpha \sin \frac{\pi(\gamma+1)}{\alpha}} \frac{1}{\lambda^{1 - \frac{\gamma+1}{\alpha}}}.$$

In particular,

$$u_\lambda(0) = \frac{1}{\alpha \sin \frac{\pi}{\alpha}} \frac{1}{\lambda^{1 - \frac{\gamma+1}{\alpha}}}.$$

(b) *For $x \in \mathbb{R}$,*

$$\int_0^\infty \frac{1 - \cos(xy)}{y^\alpha} dy = |x|^{\alpha-1} \frac{\Gamma(2 - \alpha) \sin \frac{\pi\alpha}{2}}{\alpha - 1}.$$

Proof. While the first equality in the first formula of part (a) is a consequence of [12, formula (3.241)(2)], the second equality follows from Euler's reflection formula $\Gamma(1 - z)\Gamma(z) = \frac{\pi}{\sin(\pi z)}$ which holds true for any noninteger z . The second equality of part (a) is implied by the first with $\gamma = 0$ and the formula $u_\lambda(0) = \pi^{-1} \int_0^\infty (\lambda + \theta^\alpha)^{-1} d\theta$. Part (b) follows from [34, formula (14.18)]. \square

While formula (14) of Lemma 2.2 will be used in the proof of both parts of Theorem 2.2, formula (15) will be used in the proof of Theorem 2.2(b).

Lemma 2.2. For $\alpha \in (1, 2)$ and $\lambda > 0$,

$$\lambda V_\lambda 1(x) = \mathbb{E}^x(1 - e^{-\lambda\sigma}) \sim A_{\lambda,\alpha} |x|^{\alpha-1}, \quad x \rightarrow 0, \quad (14)$$

where $A_{\lambda,\alpha} := \frac{\alpha \Gamma(2-\alpha) \sin \pi\alpha \sin \frac{\pi\alpha}{2}}{\pi(\alpha-1)} \lambda^{1-\frac{1}{\alpha}}$, and

$$\mathbb{P}^1(\sigma > y) \sim B_\alpha y^{-1+\frac{1}{\alpha}}, \quad y \rightarrow \infty, \quad (15)$$

where $B_\alpha := \frac{\sin \pi\alpha \sin \frac{\pi\alpha}{2} \Gamma(1-\alpha)}{\pi \Gamma(1-\frac{1}{\alpha})}$

Proof. We start with proving (14). According to [3, Corollary II.5.8],

$$\mathbb{E}^x e^{-\lambda\sigma} = \frac{u_\lambda(-x)}{u_\lambda(0)}, \quad x \in \mathbb{R},$$

whence

$$\begin{aligned} u_\lambda(0) \mathbb{E}^x(1 - e^{-\lambda\sigma}) &= \frac{1}{\pi} \int_0^\infty \frac{1 - \cos(x\theta)}{\lambda + |\theta|^\alpha} d\theta = \\ &= \frac{|x|^{\alpha-1}}{\pi} \int_0^\infty \frac{1 - \cos(y)}{\lambda|x|^\alpha + y^\alpha} dy \sim \frac{|x|^{\alpha-1}}{\pi} \int_0^\infty \frac{1 - \cos(y)}{y^\alpha} dy, \quad x \rightarrow 0. \end{aligned} \quad (16)$$

Invoking Lemma 2.1 we arrive at (14). Using the first equality in (16) and Lemma 2.1 we infer

$$\begin{aligned} \mathbb{E}^1(1 - e^{-\lambda\sigma}) &= \frac{\alpha \sin \frac{\pi}{\alpha} \lambda^{1-\frac{1}{\alpha}}}{\pi} \int_0^\infty \frac{1 - \cos \theta}{\lambda + \theta^\alpha} d\theta \sim \frac{\alpha \sin \frac{\pi}{\alpha}}{\pi} \int_0^\infty \frac{1 - \cos \theta}{\theta^\alpha} d\theta \lambda^{1-\frac{1}{\alpha}} \\ &= \frac{\sin \pi\alpha \sin \frac{\pi\alpha}{2} \Gamma(1-\alpha)}{\pi} \lambda^{1-\frac{1}{\alpha}}, \quad \lambda \rightarrow 0+. \end{aligned}$$

An application of Corollary 8.1.7 in [4] yields (15). \square

Lemma (2.3) is the principal ingredient of the proof of Theorem 2.2(a).

Lemma 2.3. Let $\alpha \in (1, 2)$ and $g : \mathbb{R} \rightarrow \mathbb{R}$ be a bounded measurable function satisfying $g(x) = O(|x|^{\alpha-1})$ as $x \rightarrow 0$. Assume that the function $x \mapsto \mathbb{P}(|\zeta| > x)$ is regularly varying at $+\infty$ of index $-\beta$, $\beta \in (0, \alpha - 1)$ and (2) holds. Then

$$\lim_{n \rightarrow \infty} \frac{\mathbb{E}g(\zeta/n)}{\mathbb{P}(\zeta > n)} = \int_{\mathbb{R}} g(x) \eta^*(dx) = \langle \eta^*, g \rangle < \infty,$$

where η^* is the measure defined in (4).

Proof. The functions $g^+ := \max(g, 0)$ and $g^- := \max(-g, 0)$ are nonnegative, bounded and satisfy $g^\pm(x) = O(|x|^{\alpha-1})$ as $x \rightarrow 0$. Thus, without loss of generality we can and do assume that g is nonnegative.

Put $G(x) := \mathbb{P}(\zeta \leq x)$ for $x \in \mathbb{R}$. We shall show that

$$\lim_{n \rightarrow \infty} \frac{\int_{[0, \infty)} g(x) dG(nx)}{1 - G(n)} = c_{+\beta} \int_0^\infty g(x) x^{-\beta-1} dx. \quad (17)$$

Fix any $\varepsilon \in (0, 1)$. Then

$$\lim_{n \rightarrow \infty} \frac{\int_{(\varepsilon, \infty)} g(x) dG(nx)}{1 - G(n)} = c_+ \beta \int_{\varepsilon}^{\infty} g(x) x^{-\beta-1} dx$$

follows from

$$\lim_{n \rightarrow \infty} \frac{P(\zeta > nx)}{P(\zeta > n)} = c_+ x^{-\beta}, \quad x > 0.$$

Observe that the usual requirement of continuity of g is not needed here, for the measure η^* is (absolutely) continuous.

There is a constant $c > 0$ such that $g(x) \leq cx^{\alpha-1}$ whenever $x \in (0, 1]$. With this at hand we conclude that

$$\int_{[0, \varepsilon]} g(x) dG(nx) \leq \int_{[0, \varepsilon]} cx^{\alpha-1} dG(nx) = \frac{c}{n^{\alpha-1}} \int_{[0, n\varepsilon]} x^{\alpha-1} dG(x).$$

Further,

$$\int_{[0, n\varepsilon]} x^{\alpha-1} dG(x) \sim (n\varepsilon)^{\alpha-1} (1 - G(n\varepsilon)) \frac{\beta}{\alpha - 1 - \beta} \sim \frac{\beta}{\alpha - 1 - \beta} \varepsilon^{\alpha-1-\beta} n^{\alpha-1} (1 - G(n))$$

as $n \rightarrow \infty$, where the first asymptotic relation follows from Karamata's theorem [4, Theorem 1.6.4]. We infer

$$\limsup_{n \rightarrow \infty} \int_{[0, \varepsilon]} g(x) \frac{dG(nx)}{1 - G(n)} \leq c \frac{\beta}{\alpha - 1 - \beta} \varepsilon^{\alpha-1-\beta}$$

and

$$\limsup_{n \rightarrow \infty} \int_{[0, \infty)} g(x) \frac{dG(nx)}{1 - G(n)} \leq c \frac{\beta}{\alpha - 1 - \beta} \varepsilon^{\alpha-1-\beta} + c_+ \beta \int_{\varepsilon}^{\infty} g(x) x^{-\beta-1} dx.$$

Sending $\varepsilon \rightarrow 0+$ we arrive at

$$\limsup_{n \rightarrow \infty} \int_{[0, \infty)} g(x) \frac{dG(nx)}{1 - G(n)} \leq c_+ \beta \int_0^{\infty} g(x) x^{-\beta-1} dx.$$

For the lower bound, write, for any $\varepsilon > 0$,

$$\int_{[0, \infty)} g(x) \frac{dG(nx)}{1 - G(n)} \geq \int_{(\varepsilon, \infty)} g(x) \frac{dG(nx)}{1 - G(n)} \rightarrow c_+ \beta \int_{\varepsilon}^{\infty} g(x) x^{-\beta-1} dx, \quad n \rightarrow \infty.$$

Sending $\varepsilon \rightarrow 0+$ we obtain

$$\liminf_{n \rightarrow \infty} \int_{[0, \infty)} g(x) \frac{dG(nx)}{1 - G(n)} \geq c_+ \beta \int_0^{\infty} g(x) x^{-\beta-1} dx,$$

and (17) follows.

Starting with

$$\lim_{n \rightarrow \infty} \frac{P(-\zeta > nx)}{P(\zeta > n)} = c_- x^{-\beta}, \quad x > 0$$

and arguing analogously we also conclude that

$$\int_{(-\infty, 0)} g(x) \frac{dG(nx)}{1 - G(n)} = c_- \beta \int_{-\infty}^0 g(x) |x|^{-\beta-1} dx.$$

Combining this with (17) completes the proof of the lemma. \square

AUXILIARY RESULTS FOR THE CASE $\beta > \alpha - 1$ OR $E|\zeta| < \infty$. Lemmas 2.4 and 2.5 will be used for the proof of Theorem 2.2 (b).

Lemma 2.4. For $\lambda > 0$ and any bounded and continuous function $f : \mathbb{R} \rightarrow \mathbb{R}$,

$$\lim_{x \rightarrow 0} \frac{V_\lambda f(x)}{V_\lambda 1(x)} = \lambda R_\lambda^{U_\alpha} f(0) = \lambda \int_{\mathbb{R}} u_\lambda(y) f(y) dy. \quad (18)$$

Remark 2.1. According to this lemma and [5, Theorem 4.2, Section V.4] there exists a unique recurrent extension of the process U_α killed at 0 which has a zero sojourn at 0. As a consequence, this extension coincides with U_α . More details on the excursion theory, recurrent extensions of Markov processes etc. will be given in Section 3.

Proof. For each $x \in \mathbb{R}$, denote by X_x the process defined in (1) in which we formally replace ζ with x . For $k \in \mathbb{N}$, denote by $\sigma_k^{(x)}$ the time of the k th jump (of size x) from 0 and note that the random variables $\sigma_1^{(x)}, \sigma_2^{(x)}, \dots$ are independent and identically distributed. We shall use a representation

$$X_x(t) = U_\alpha(t) + xN_\alpha(x, t), \quad t \geq 0, x \in \mathbb{R},$$

where $N_\alpha(x, t) = \#\{k \in \mathbb{N} : \sigma_1^{(x)} + \dots + \sigma_k^{(x)} \leq t\}$.

Arguing as in (9) and (10) we conclude that

$$\frac{V_\lambda f(x)}{\lambda V_\lambda 1(x)} = E \int_0^\infty e^{-\lambda t} f(X_x(t)) dt, \quad x \in \mathbb{R}.$$

In view of this and the Lebesgue dominated convergence theorem, we are left with showing that

$$xN_\alpha(x, t) \xrightarrow{P} 0, \quad x \rightarrow 0. \quad (19)$$

To prove (19), recall that the process U_α is self-similar of index $1/\alpha$. This implies that, for $k \in \mathbb{N}$, $\sigma_k^{(x)}$ has the same distribution as $\sigma_k^{(1)}|x|^\alpha$ and thereupon $N_\alpha(x, t)$ has the same distribution as $\#\{k \in \mathbb{N} : \sigma_1^{(1)} + \dots + \sigma_k^{(1)} \leq t|x|^{-\alpha}\}$. By [22, Theorem 3.2],

$$P^1(\sigma > t|x|^{-\alpha}) N_\alpha(x, t) \xrightarrow{d} S_{1-\frac{1}{\alpha}}^{\leftarrow}(1) := \inf\{t \geq 0 : S_{1-\frac{1}{\alpha}}(t) > 1\}, \quad x \rightarrow 0,$$

where \xrightarrow{d} denotes convergence of one-dimensional distributions, and $(S_{1-\frac{1}{\alpha}}(t))_{t \geq 0}$ is a drift-free $(1 - \frac{1}{\alpha})$ -stable subordinator with

$$-\log Ee^{-uS_{1-\frac{1}{\alpha}}(1)} = \Gamma(1/\alpha)u^{1-\frac{1}{\alpha}}, \quad u \geq 0.$$

Since, by (15), $P^1(\sigma > t|x|^{-\alpha}) \sim B_\alpha t^{-1+\frac{1}{\alpha}}|x|^{\alpha-1}$ as $x \rightarrow 0$, we infer (19). \square

Lemma 2.5. Let $a : \mathbb{R} \rightarrow [0, \infty)$ and $b : \mathbb{R} \rightarrow \mathbb{R}$ be functions such that $|b(x)| \leq Ca(x)$ for all $x \in \mathbb{R}$ and some constant $C > 0$ and that $\lim_{x \rightarrow 0} \frac{b(x)}{a(x)} = A \in \mathbb{R}$. Assume that a sequence $(\mu_n)_{n \geq 1}$ of probability measures satisfies $\mu_n(\{0\}) = 0$ for large n and, for each $\delta > 0$,

$$\lim_{n \rightarrow \infty} \frac{\int_{|x| \geq \delta} a(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)} = 0. \quad (20)$$

Then

$$\lim_{n \rightarrow \infty} \frac{\int_{\mathbb{R}} b(x) \mu_n(dx)}{\int_{\mathbb{R}} a(x) \mu_n(dx)} = A.$$

Proof. Let $\varepsilon > 0$ be arbitrary. Select $\delta > 0$ such that $|\frac{b(x)}{a(x)} - A| < \varepsilon$ whenever $|x| \leq \delta$. Then

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{\int_{\mathbb{R}} b(x) \mu_n(dx)}{\int_{\mathbb{R}} a(x) \mu_n(dx)} &= \limsup_{n \rightarrow \infty} \frac{\frac{\int_{|x| \geq \delta} b(x) \mu_n(dx) + \int_{|x| < \delta} b(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)}}{\frac{\int_{|x| \geq \delta} a(x) \mu_n(dx) + \int_{|x| < \delta} a(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)}} = \limsup_{n \rightarrow \infty} \frac{\frac{\int_{|x| < \delta} b(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)}}{\frac{\int_{|x| < \delta} a(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)}} = \\ &= \limsup_{n \rightarrow \infty} \frac{\int_{|x| < \delta} b(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)} \leq \limsup_{n \rightarrow \infty} \frac{(A + \varepsilon) \int_{|x| < \delta} a(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)} = A + \varepsilon \end{aligned}$$

having utilized for the second equality

$$\lim_{n \rightarrow \infty} \frac{\int_{|x| \geq \delta} b(x) \mu_n(dx)}{\int_{|x| < \delta} a(x) \mu_n(dx)} = 0$$

which is a consequence of (20) and the assumption $|b(x)| \leq Ca(x)$ for all $x \in \mathbb{R}$. An analogous inequality for the lower limit follows similarly. Sending $\varepsilon \rightarrow 0$ completes the proof of the lemma. \square

2.3. Proof of Theorem 2.2. We first prove the theorem in the case $\beta < \alpha - 1$ (part (a) of the theorem). In view of (11) we have to show that

$$\lim_{n \rightarrow \infty} \frac{f(0)/m_n + \langle P_{\zeta/n}, V_{\lambda} f \rangle}{1/m_n + \langle P_{\zeta/n}, V_{\lambda} 1 \rangle} = \frac{\langle \eta^*, V_{\lambda} f \rangle}{\langle \eta^*, V_{\lambda} 1 \rangle}. \quad (21)$$

The function $V_{\lambda} 1$ is nonnegative and bounded (by λ^{-1}). According to (14), it satisfies $V_{\lambda} 1(x) = O(|x|^{\alpha-1})$ as $x \rightarrow 0$. By virtue of

$$|V_{\lambda} f(x)| \leq \sup_{y \in \mathbb{R}} |f(y)| V_{\lambda} 1(x), \quad x \in \mathbb{R} \quad (22)$$

and the fact that f is bounded by assumption, we conclude that $V_{\lambda} f$ is a bounded function satisfying $V_{\lambda} f(x) = O(|x|^{\alpha-1})$ as $x \rightarrow 0$. Hence, an application of Lemma (2.3) yields

$$\langle P_{\zeta/n}, V_{\lambda} f \rangle \sim P(\zeta > n) \langle \eta^*, V_{\lambda} f \rangle \quad \text{and} \quad \langle P_{\zeta/n}, V_{\lambda} 1 \rangle \sim P(\zeta > n) \langle \eta^*, V_{\lambda} 1 \rangle, \quad n \rightarrow \infty.$$

These together with our choice of m_n prove (21).

Now we are turning to the proof of part (b). Formula (11) tells us that we have to prove that

$$\lim_{n \rightarrow \infty} \frac{f(0)/m_n + \langle P_{\zeta/n}, V_{\lambda} f \rangle}{1/m_n + \langle P_{\zeta/n}, V_{\lambda} 1 \rangle} = \lambda R_{\lambda}^{U_{\alpha}} f(0). \quad (23)$$

In the notation of Lemma 2.5, put $a := V_{\lambda} 1$, $b := V_{\lambda} f$ and $\mu_n := P_{\zeta/n}$ for $n \geq 1$. Then, in view of (22), we may put $C := \sup_{y \in \mathbb{R}} |f(y)|$ and, by Lemma 2.4, $A := \lambda R_{\lambda}^{U_{\alpha}} f(0)$. Now relation (23) follows from Lemma 2.5 and our choice of m_n if we can show that

$$\lim_{n \rightarrow \infty} m_n \langle P_{\zeta/n}, V_{\lambda} 1 \rangle = \infty \quad (24)$$

and that, for any $\delta > 0$,

$$\lim_{n \rightarrow \infty} \frac{\int_{|x| \geq \delta} \mathbb{E}^x(1 - e^{-\lambda\sigma}) P_{\zeta/n}(dx)}{\int_{|x| < \delta} \mathbb{E}^x(1 - e^{-\lambda\sigma}) P_{\zeta/n}(dx)} = 0 \quad (25)$$

which is condition (20) with $a(x) = V_{\lambda} 1(x) = \lambda^{-1} \mathbb{E}^x(1 - e^{-\lambda\sigma})$ for $x \in \mathbb{R}$.

Fix any $\delta > 0$. On the one hand,

$$\int_{|x| \geq \delta} \mathbb{E}^x(1 - e^{-\lambda\sigma}) P_{\zeta/n}(dx) \leq \mathbb{P}(|\zeta| \geq n\delta) = o(n^{1-\alpha}), \quad n \rightarrow \infty$$

which holds true if $\mathbb{E}|\zeta| < \infty$ or $\beta > \alpha - 1$. On the other hand, appealing to (14) we conclude that, for appropriate constant $K > 0$,

$$\begin{aligned} \int_{|x| < \delta} \mathbb{E}^x(1 - e^{-\lambda\sigma}) P_{\zeta/n}(dx) &\geq K \int_{|x| < \delta} |x|^{\alpha-1} P_{\zeta/n}(x) = Kn^{1-\alpha} \int_{|x| < n\delta} |x|^{\alpha-1} P_{\zeta}(dx) \\ &\sim K\mathbb{E}(|\zeta|^{\alpha-1})n^{1-\alpha}, \quad n \rightarrow \infty. \end{aligned} \quad (26)$$

Since $\mathbb{E}(|\zeta|^{\alpha-1}) < \infty$ if $\mathbb{E}|\zeta| < \infty$ or $\beta > \alpha - 1$, (25) follows. Finally, (24) is a consequence of (26) and our assumption $\lim_{n \rightarrow \infty} m_n n^{1-\alpha} = \infty$.

The proof of Theorem 2.2 is complete.

3. ITÔ'S EXCURSION THEORY

According to Corollary 2.1, the processes $X_{\zeta/n}$ converge in distribution. If $\beta > \alpha - 1$ or $\mathbb{E}|\zeta| < \infty$, the limit process is U_{α} and, as such, well-understood. If $\beta < \alpha - 1$, we only have a resolvent description of the limit process that we call a skew α -stable Lévy process. In this section, we give a probabilistic representation of this process via Itô's excursion theory and an SDE, which include a local time of the process. There are several definitions of a local time of a Markov process. Usually these produce the same process, up to a multiplicative constant. To provide an absolutely rigorous formulation of our results, we should stick to a particular definition that serves our needs. As far as the skew stable Lévy process is concerned, the exact values of constants are as important as the parameter of permeability appearing in the defining equation (5) for the SBM.

We briefly review below some basic facts of Itô's excursion theory for Markov processes. We only discuss real-valued processes and leave aside processes taking values in the other spaces. We follow Blumenthal's book [5] and cite specific theorems or point out pages for the most important results. Let (X, P^x) , $x \in \mathbb{R}$ be a Feller process, that is, a Markov process whose transition semigroup $(P_t)_{t \geq 0}$ is strongly continuous on $C_0(\mathbb{R})$. Without loss of generality we always assume that all processes to be considered are standard, see the corresponding definition in [6, Chapter I, §9]. Assume that X is recurrent at 0, that is, $P^x(\sigma < \infty) = 1$ for all $x \in \mathbb{R}$, where, as before, $\sigma = \sigma(X)$ denotes the first hitting time of 0.

Associated with (X, P^x) are

- (i) $\bar{X}(t) := X(t \wedge \sigma)$, $t \geq 0$ the process stopped at 0 (we denote by \bar{P}^x and \bar{P}_t its distribution and semigroup, respectively), and
- (ii) the process killed at 0 with the semigroup P_t^0 and the transition probabilities $P^0(t, x, A) = P(X(t) \in A, t < \sigma)$ for $t \geq 0$ and $x \in \mathbb{R}$.

The killed process and its semigroup will be called *minimal process* and *minimal semigroup*, respectively.

Below we recall elements of Itô's synthesis theory, which describes all recurrent extensions of the minimal process.

Assume that

the function $x \mapsto \bar{\mathbb{E}}^x e^{-\lambda\sigma}$ is continuous for each $\lambda > 0$ and $\lim_{|x| \rightarrow \infty} \bar{\mathbb{E}}^x e^{-\sigma} = 0$.

Remark 3.1. All the properties of a minimal semigroup, stated above and below, are satisfied for the minimal semigroup of a symmetric α -stable Lévy process with $\alpha \in (1, 2)$.

To each function $u \in D([0, \infty))$, we associate $\sigma(u) := \inf\{t > 0 : u(t) = 0\}$. Let \hat{P} be a σ -finite measure on D supported by the set of functions $\{u : u(t) = 0, t \geq \sigma(u)\}$. The elements of this set will be called *excursions* and $\sigma(u)$ will be called the *length of excursion* u . We shall assume that

$$\hat{P}(1 - e^{-\sigma}) \leq 1$$

and that $\hat{P}(|u(0)| > x) < \infty$ for $x > 0$.

Let $N(ds, du)$ denote a Poisson point measure on $[0, \infty) \times D$ with intensity $ds \times \hat{P}(du)$. Denote by (s_k, u_k) the atoms of N , that is, $N = \sum_k \delta_{(s_k, u_k)}$. Put $m := 1 - \hat{P}(1 - e^{-\sigma})$,

$$\tau(s) := ms + \sum_{s_k \leq s} \sigma(u_k) = ms + \int_{[0, s]} \int_D \sigma(u) N(dz, du), \quad s \geq 0$$

and

$$\varphi(t) := \inf\{s \geq 0 : \tau(s) > t\}, \quad t \geq 0. \quad (27)$$

Assume that at least one of the following conditions holds:

$$m > 0 \text{ or the measure } \hat{P} \text{ is infinite.}$$

Then $(\tau(s))_{s \geq 0}$ is a strictly increasing subordinator. For $t \in [\tau(s_k -), \tau(s_k)]$, put $X(t) := u_k(t - \tau(s_k -))$ and, for $t \notin \cup_k [\tau(s_k -), \tau(s_k)]$, put $X(t) := 0$.

Further, we assume that the characteristic measure \hat{P} is *compatible with the minimal semigroup*. This means that, for all $n \geq 1$, all $0 \leq s_1 \leq \dots \leq s_n \leq s$ and all bounded measurable functions g, g_1, \dots, g_n with $g(0) = 0$,

$$\hat{P}\left(g(u(t+s))g_1(u(s_1)) \dots g_n(u(s_n)); \sigma > s\right) = \hat{P}\left(P_t^0 g(u(s))g_1(u(s_1)) \dots g_n(u(s_n)); \sigma > s\right),$$

and

$$\hat{P}\left(g(u(t)); u(0) \in B\right) = \hat{P}\left(P_t^0 g(u(0)); u(0) \in B\right),$$

(see [5, Chapter V §1 and V §2 (d)] for more details). Then, according to [5, Theorem 2.10, Chapter V §2]), X is a Markov extension of the minimal process with $X(0) = 0$. Moreover, each extension (starting at 0) can be obtained with the help of merging excursions, under suitable measure \hat{P} satisfying the compatibility conditions. An extension has a zero sojourn at 0 if, and only if, $m = 0$.

Remark 3.2. If $X(0) \neq 0$, then the behavior of X until hitting 0 is determined uniquely by the minimal semigroup or the semigroup corresponding to the stopped process.

Note that the measure \hat{P} is finite if, and only if, $\hat{P} = \bar{P}^\theta$ for some finite measure θ on $\mathbb{R} \setminus \{0\}$. In this case, X is a holding and jumping process, with holding time $\|\theta\|$ and jumping distribution $\frac{\theta}{\|\theta\|}$, where $\|\theta\| = \theta(\mathbb{R} \setminus \{0\})$.

If the measure \hat{P} is infinite, then X can be obtained as an a.s. limit of holding and jumping processes. To explain the construction, recall that [5, Chapter 5.2] a family of σ -finite measures $(\theta_t)_{t>0}$ on the Borel subsets of $\mathbb{R} \setminus \{0\}$ is called *entrance law* for P_t^0 , if

$$\theta_s P_t^0 = \theta_{t+s}, \quad t \geq 0, \quad s > 0$$

and

$$\langle \theta_s, V1 \rangle = E^{\theta_s}(1 - e^{-\sigma}) \leq 1, \quad s > 0.$$

If θ is a σ -finite measure on $\mathbb{R} \setminus \{0\}$ such that $E^\theta(1 - e^{-\sigma}) \leq 1$, then $\theta_t := \theta P_t^0$ is an entrance law. Another important example of an entrance law is $\hat{\theta}_t(A) := \hat{P}(X_t \in A, t < \sigma)$.

Consider a family of holding and jumping processes $(X_\varepsilon)_{\varepsilon>0}$ with jumping distribution $\frac{\hat{\eta}_\varepsilon}{\|\hat{\eta}_\varepsilon\|}$ and holding parameter $m + \delta_\varepsilon$, where $\delta_\varepsilon > 0$ is any function satisfying $\delta_\varepsilon = o(\varepsilon)$ as $\varepsilon \rightarrow 0$. It follows from the proof of [5, Theorem 2.10, Chapter 5.2] that the corresponding resolvents converge uniformly. By Theorem 2.1, this entails the distributional convergence $X_\varepsilon \Rightarrow X$ on D as $\varepsilon \rightarrow 0$. The aforementioned proof contains an explicit construction of X_ε , which includes merging of excursions of the length larger than $\varepsilon > 0$, and such that $\lim_{\varepsilon \rightarrow 0} X_\varepsilon(t) = X(t)$ a.s.

Note also that, for any entrance law (θ_t) , there exists a unique characteristic measure \hat{P} satisfying $\theta_t = \hat{\theta}_t, t > 0$, see [5, Theorem 4.7, Chapter V]. Moreover, the characteristic measure \hat{P} can be recovered from the process X , see [5, Chapter III], and the process $(\varphi(t))_{t \geq 0}$ defined in (27) is the Blumenthal-Gettoor local time of X at 0, see [5, Theorem 2.3, Chapter V §2], that is, a continuous additive functional which satisfies

$$E^x e^{-\sigma} = E^x \int_0^\infty e^{-t} d\varphi(t), \quad x \in \mathbb{R}.$$

Existence, uniqueness and some other properties of the local time are discussed in [5, p. 91-93]. Therefore, under some natural properties imposed on the minimal semigroup, there is a one-to-one correspondence between the recurrent extension of the minimal process, the characteristic measure \hat{P} and the entrance law (θ_t) . We stress that the analysis of the entrance law is simpler than that of the two other objects.

Any entrance law for P_t^0 can be uniquely decomposed as the sum of two entrance laws

$$\theta_t = \rho_t + \theta P_t^0, \quad t > 0,$$

where θ is a σ -finite measure, and the measure ρ_t satisfies

$$\lim_{t \rightarrow 0^+} \rho_t(\mathbb{R} \setminus [-x, x]) = 0$$

for any $x > 0$, see [5, p. 140]. These entrance laws, which are called *continuous entrance* and *jump entrance*, respectively, can also be characterized as follows:

$$\theta_t^c(\cdot) := \rho_t = \hat{P}(X(s) \in \cdot, X(0) = 0, s < \sigma),$$

$$\theta_t^j(\cdot) := \hat{P}(X(s) \in \cdot, X(0) \neq 0, s < \sigma),$$

see [5, p. 156]. It follows from Theorems 2.6, 2.8 and the proof of Theorem 2.10 from [5, Chapter V] that if an entrance law is a jump entrance law, that is, $\theta_t = \theta P_t^0$, where θ is an infinite measure satisfying $E^\theta(1 - e^{-\sigma}) = 1$, then the corresponding recurrent extension X has a zero sojourn at 0, its resolvent satisfies

$$\lambda R_\lambda f(0) = \frac{\langle \theta, V_\lambda f \rangle}{\langle \theta, V_\lambda 1 \rangle}, \quad (28)$$

the corresponding characteristic measure \hat{P} is equal to \bar{P}^θ .

Here is another conclusion motivated by the aforementioned facts. Assume that there exists a unique continuous entrance law ρ_t with the characteristic measure \hat{P}^c that satisfies $\hat{P}^c(1 - e^{-\sigma}) = 1$. For instance, this is the case for the process U_α , see Remark 2.1. Then any entrance law (θ_t) , which corresponds to an extension with a zero sojourn at 0, can be uniquely represented by

$$\theta_t = (1 - p)\rho_t + p\theta P_t^0, \quad (29)$$

where $p \in [0, 1]$, and the measure θ satisfies $\mathbb{E}^\theta(1 - e^{-\sigma}) = 1$. The corresponding characteristic measure is

$$\hat{P} = (1 - p)\hat{P}^c + p\bar{P}^\theta. \quad (30)$$

4. SDE FOR THE SKEW STABLE LÉVY PROCESS

In this section all semigroups, resolvents, etc. are related to extensions of the minimal process corresponding to the process U_α killed at 0.

4.1. Existence of solutions to SDEs with local times. Let X be a Feller process with the resolvent

$$R_\lambda f(x) = V_\lambda f(x) + \mathbb{E}^x e^{-\lambda\sigma} \frac{\langle \theta, V_\lambda f \rangle}{\langle \theta, V_\lambda 1 \rangle}, \quad \lambda > 0, x \in \mathbb{R}, \quad (31)$$

where θ is an infinite measure satisfying $\mathbb{E}^\theta(1 - e^{-\sigma}) = 1$. The right-hand side in (31) is obtained by a combination of (9) and (28).

Theorem 4.1. *The process X is a (weak) solution to the SDE*

$$X(t) = X(0) + U_\alpha(t) + S_\theta(L_0^X(t)), \quad t \geq 0, \quad (32)$$

where U_α is a symmetric α -stable Lévy process, S_θ is a pure-jump Lévy process with the Lévy measure θ , which is independent of U_α , the process L_0^X is the local time of X at 0 (see the definition in the previous subsection).

A specialization of Theorem 4.1, with the Lévy measure $\theta = \eta$ given by (7) and $S_\theta = S_\beta$, immediately proves Theorem B, except the explicit formula for the constant C . Here is the remaining piece of the proof. Write

$$\begin{aligned} \frac{1}{C} &= \int_{\mathbb{R}} \mathbb{E}^x(1 - e^{-\sigma(U_\alpha)})\eta(dx) = \frac{c_- + c_+}{\pi u_1(0)} \int_0^\infty \int_0^\infty \frac{1 - \cos(x\theta)}{1 + \theta^\alpha} d\theta \frac{dx}{x^{1+\beta}} \\ &= \frac{c_- + c_+}{\pi u_1(0)} \frac{\Gamma(1 - \beta) \sin \frac{\pi(1+\beta)}{2}}{\beta} \int_0^\infty \frac{\theta^\beta}{1 + \theta^\alpha} d\theta = \frac{(c_- + c_+)\Gamma(1 - \beta) \cos \frac{\pi\beta}{2} \sin \frac{\pi}{\alpha}}{\beta \sin \frac{\pi(\beta+1)}{\alpha}} \end{aligned}$$

having utilized (7) and (16) for the second equality, Fubini's theorem and Lemma 2.1(b) for the third equality and Lemma 2.1(a) for the last equality.

Proof of Theorem 4.1. Since $\mathbb{E}^\theta(1 - e^{-\sigma}) = 1$ by assumption, we conclude that the 'holding' parameter $m = 1 - \mathbb{E}^\theta(1 - e^{-\sigma})$ is equal to 0, and the process X has a zero sojourn at 0 a.s.

Let $N_\theta(ds, dx)$ be a Poisson point measure with intensity $ds \times \theta(dx)$. Let $((s_k, x_k))_{k \geq 1}$ be a (measurable) enumeration of the atoms of N_θ . Let $U_{\alpha,1}, U_{\alpha,2}, \dots$ denote independent copies of U_α , which are independent of N_θ . Put $N := \sum_{k \geq 1} \delta_{(s_k, x_k + U_{\alpha,k}(\cdot \wedge \sigma_k))}$, where

$$\sigma_k := \sigma(x_k + U_{\alpha,k}(\cdot)) := \inf\{t \geq 0 : x_k + U_{\alpha,k}(t) = 0\}.$$

Then N is a Poisson random measure on $[0, \infty) \times D$ with intensity $ds \times \hat{P} = ds \times \bar{P}^\theta$.

Without loss of generality we can and do assume that $X(0) = 0$ and that the process X is built upon the Poisson point measure N with the help of Itô's procedure. Here are details of the construction. Put

$$\tau(s) := \sum_{s_k \leq s} \sigma(x_k + U_{\alpha,k}(\cdot)), \quad s \geq 0,$$

and

$$\varphi(t) := \inf\{s \geq 0 : \tau(s) > t\}, \quad t \geq 0.$$

For $t \in [\tau(s_k-), \tau(s_k))$, put $X(t) := x_k + U_{\alpha,k}(t - \tau(s_k-))$, and, for $t \notin \bigcup_k [\tau(s_k-), \tau(s_k))$, put $X(t) := 0$.

Lemma 4.1. *For almost all $t > 0$ with respect to the Lebesgue measure,*

$$\mathbb{P}\left(\exists k : t \in (\tau(s_k-), \tau(s_k)) \text{ and } \tau(s_k-) = \varphi(t)\right) = 1.$$

The proof follows from the fact that X has a zero sojourn at 0 and Fubini's theorem.

Put

$$N_\theta^{(\varepsilon)} := \sum_k \mathbb{I}_{\{|x_k| > \varepsilon\}} \delta_{(s_k, x_k)}, \quad N^{(\varepsilon)} := \sum_{k \geq 1} \mathbb{I}_{\{|x_k| > \varepsilon\}} \delta_{(s_k, x_k + U_{\alpha,k}(\cdot \wedge \sigma_k))}. \quad (33)$$

Then $N_\theta^{(\varepsilon)}$ is a Poisson point measure on $[0, \infty) \times (\mathbb{R} \setminus \{0\})$ with intensity $ds \times (\mathbb{I}_{\{|x| > \varepsilon\}} \theta(dx))$, and $N^{(\varepsilon)}$ is a Poisson point measure on $[0, \infty) \times D$ with intensity $ds \times \bar{P} \mathbb{I}_{\{|x| > \varepsilon\}} \theta(dx)$.

Assuming that $m = 0$ we now construct a process $X^{(\varepsilon)}$ with the help of the Poisson random measure $N^{(\varepsilon)}$, along the lines of construction of X with the help of N . Put

$$\tau^{(\varepsilon)}(s) := \sum_{s_k \leq s, |x_k| > \varepsilon} \sigma(x_k + U_{\alpha,k}(\cdot)), \quad s \geq 0$$

and

$$\varphi^{(\varepsilon)}(t) := \inf\{s \geq 0 : \tau^{(\varepsilon)}(s) \geq t\}, \quad t \geq 0.$$

For $t \in [\tau^{(\varepsilon)}(s_k-), \tau^{(\varepsilon)}(s_k))$, put $X^{(\varepsilon)}(t) := x_k + U_{\alpha,k}(t - \tau^{(\varepsilon)}(s_k-))$ provided that $|x_k| > \varepsilon$. Otherwise, we put $X^{(\varepsilon)}(t) := 0$.

Remark 4.1. In contrast to the process X , a.s. there does not exist a t such that $X^{(\varepsilon)}(t) = 0$. The process $X^{(\varepsilon)}$ jumps upon ‘touching’ 0.

In view of $\int_{\mathbb{R}} \int_0^T \mathbb{I}_{\{|x| > \varepsilon\}} ds \eta(dx) < \infty$ for any $T > 0$, the Poisson point process $N_\eta^{(\varepsilon)}$ can be represented by $\sum_{k \geq 1} \delta_{(s_k^{(\varepsilon)}, x_k^{(\varepsilon)})}$, where $0 < s_1^{(\varepsilon)} < s_2^{(\varepsilon)} < s_3^{(\varepsilon)} < \dots$, and each interval $[0, T]$ contains finitely many $s_k^{(\varepsilon)}$ a.s. Put $\sigma_k^{(\varepsilon)} := \sigma(x_k^{(\varepsilon)} + U_{\alpha,k}^{(\varepsilon)}(\cdot))$, where $U_{\alpha,k}^{(\varepsilon)}$ is a process from the collection $(U_{\alpha,j})$, which corresponds to $s_k^{(\varepsilon)}$ in the representation of N . According to the construction of $X^{(\varepsilon)}$,

$$X^{(\varepsilon)}(0) = x_1^{(\varepsilon)},$$

$$X^{(\varepsilon)}(\sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)}) = x_{k+1}^{(\varepsilon)},$$

$$X^{(\varepsilon)}(t) = x_k^{(\varepsilon)} + U_{\alpha,k}^{(\varepsilon)}(t), \quad t \in [\sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)}, \sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)} + \sigma_{k+1}^{(\varepsilon)}),$$

and

$$\sigma_{k+1}^{(\varepsilon)} = \inf\{t \geq \sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)} : x_k^{(\varepsilon)} + U_{\alpha,k}^{(\varepsilon)}(t - (\sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)})) = 0\}.$$

Plainly, the process $U_\alpha^{(\varepsilon)}$ defined by $U_\alpha^{(\varepsilon)}(0) := 0$, and

$$U_\alpha^{(\varepsilon)}(t) := U_\alpha^{(\varepsilon)}(\sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)}) + U_{\alpha,k}^{(\varepsilon)}(t - (\sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)})),$$

for $t \in [\sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)}, \sigma_1^{(\varepsilon)} + \dots + \sigma_k^{(\varepsilon)} + \sigma_{k+1}^{(\varepsilon)})$, is a symmetric α -stable Lévy process, which is independent of $N_\theta^{(\varepsilon)}$. In particular, $U_\alpha^{(\varepsilon)}$ is independent of

$$S_\theta^{(\varepsilon)}(s) := \int_{[0, s]} \int_{|x| > \varepsilon} x N_\theta^{(\varepsilon)}(dz, dx) = \int_{[0, s]} \int_{\mathbb{R} \setminus \{0\}} x \mathbb{I}_{\{|x| > \varepsilon\}} N_\theta(dz, dx). \quad (34)$$

Note also that

$$X^{(\varepsilon)}(t) = U_\alpha^{(\varepsilon)}(t) + S_\theta^{(\varepsilon)}(\varphi^{(\varepsilon)}(t)), \quad t \geq 0.$$

Remark 4.2. Let ζ_ε be a random variable with distribution $\mathbb{P}(\zeta_\varepsilon \in A) = \frac{\int_{|x| > \varepsilon} \mathbb{I}_A(x) \theta(dx)}{\theta(|x| > \varepsilon)}$. Then the distribution of X_ε coincides with that of X_{ζ_ε} , see (1) for the definition.

Lemma 4.2. *For each $T > 0$, almost surely*

$$\lim_{\varepsilon \rightarrow 0} \sup_{s \in [0, T]} |\tau^{(\varepsilon)}(s) - \tau(s)| = 0,$$

$$\lim_{\varepsilon \rightarrow 0} \sup_{t \in [0, T]} |\varphi^{(\varepsilon)}(t) - \varphi(t)| = 0$$

and

$$\lim_{\varepsilon \rightarrow 0} \sup_{s \in [0, T]} |S_\theta^{(\varepsilon)}(s) - S_\theta(s)| = 0,$$

where $S_\theta(s) = \int_0^s \int_{\mathbb{R} \setminus \{0\}} x N_\theta(dz, dx)$ for $s \geq 0$.

The proof follows from the construction and the fact that the processes $(\tau(t))$ and S_θ are pure-jump Lévy processes of (locally) finite variation.

Lemma 4.3. *For almost all $t > 0$ with respect to the Lebesgue measure,*

$$\mathbb{P}(\exists \varepsilon_0 > 0 \forall \varepsilon \in (0, \varepsilon_0) : \varphi^{(\varepsilon)}(t) = \varphi(t)) = 1. \quad (35)$$

Proof. It follows from Lemmas 4.1 and 4.2 that for a.e. $t > 0$ with probability 1 there exists k such that

$$t \in (\tau(s_{k-}), \tau(s_k)) \quad \text{and} \quad \tau(s_{k-}) = \varphi(t)$$

and

$$\lim_{\varepsilon \rightarrow 0} \tau^{(\varepsilon)}(s_{k-}) = \tau(s_{k-}), \quad \lim_{\varepsilon \rightarrow 0} \tau^{(\varepsilon)}(s_k) = \tau(s_k).$$

These entail $t \in (\tau^{(\varepsilon)}(s_{k-}), \tau^{(\varepsilon)}(s_k))$ for small $\varepsilon > 0$ and thereupon (35). \square

Corollary 4.1. *For each $t > 0$,*

$$\lim_{\varepsilon \rightarrow 0} S_\theta^{(\varepsilon)}(\varphi^{(\varepsilon)}(t)) = S_\theta(\varphi(t)) \quad a.s.$$

A fixed time t_0 is not a jump-time of a Lévy process a.s. Hence, by Fubini's theorem, for each $k \geq 1$ and almost all $t_0 > 0$ with respect to the Lebesgue measure,

$$\mathbb{P} \left(\lim_{t \rightarrow t_0} U_{\alpha, k}(t) = U_{\alpha, k}(t_0) \right) = 1 \quad a.s.$$

This together with Lemmas 4.1 and 4.2 enables us to conclude that

$$\mathbb{P} \left(\lim_{t \rightarrow t_0} X^{(\varepsilon)}(t) = X(t_0) \right) = 1$$

for almost all $t_0 > 0$. Invoking Corollary 4.1 we infer

$$\lim_{t \rightarrow t_0} U_\alpha^{(\varepsilon)}(t) = \lim_{t \rightarrow t_0} \left(X^{(\varepsilon)}(t) - S_\theta^{(\varepsilon)}(\varphi^{(\varepsilon)}(t)) \right) = X(t_0) - S_\theta(\varphi(t_0)) \quad \text{a.s.} \quad (36)$$

for almost all $t_0 > 0$.

Denote the process $(X(t) - S_\theta(\varphi(t)))$ by U_α . The paths of this process are càdlàg, for so are the paths of X and $(S_\theta(\varphi(t)))$. The process U_α is a symmetric α -stable Lévy process as a limit of symmetric α -stable Lévy processes. Hence, (36) entails

$$X(t) = U_\alpha(t) + S_\theta(\varphi(t)), \quad t \geq 0 \quad \text{a.s.}$$

Since $(U_\alpha^{(\varepsilon)}(t))$ is independent of S_θ , so is U_α . □

Remark 4.3. We have not proved that $\lim_{\varepsilon \rightarrow 0} X^{(\varepsilon)} = X$ a.s. on D . However, this has never been claimed.

Remark 4.4. For each $T > 0$, the process $(U_\alpha^{(\varepsilon)}(t))_{t \in [0, T]}$ was obtained by merging finitely (but randomly) many fragments of paths of independent α -stable Lévy processes. However, the process $(U_\alpha(t))_{t \in [0, T]}$ is built upon a countable number of paths. We cannot offer a good formula for $U_\alpha(t)$, other than $U_\alpha(t) = X(t) - S_\theta(\varphi(t))$ and attract the reader attention to the fact that the set $\{t \geq 0 : X(t) = 0\}$ is uncountable (the set does not coincide with a countable set of the endpoints of excursion intervals). A similar noise representation, in terms of the difference of a Markov process and a generalized drift, is typical in the framework of diffusions with semipermeable membrane, see [31, 33]. In this context the Markov process Y , say is first constructed via a semigroup technique. Then the generalized drift is identified with an additive functional of Y . Finally, it has to be checked that the difference of Y and the generalized drift is a stochastic integral.

The Lévy measure θ of the process S_θ appearing in (32) satisfies $E^\theta(1 - e^{-\sigma}) = 1$. Below we construct a process, similar to S_θ , which satisfies (32) when the latter equality does not necessarily holds. Denote by (ρ_t) , \hat{P}_α and N_α the entrance law, the intensity of excursion measure and the corresponding Poisson point measure of a symmetric α -stable Lévy process. Recall (see Lemma 2.4 and Remark 2.1) that there exists a unique continuous entrance law with a zero sojourn at 0, and that this entrance law corresponds to a symmetric α -stable Lévy process. Let θ be a sigma-finite measure on $\mathbb{R} \setminus \{0\}$ satisfying $E^\theta(1 - e^{-\sigma}) = 1$. Define the entrance law $\theta_t := \theta P_t^0$, the characteristic measure \bar{P}^θ and the corresponding Poisson point measure N , which is independent of N_α . For $p \in [0, 1]$, define the entrance law $p\theta_t + (1 - p)\rho_t$, the intensity of excursions measure $p\bar{P}^\theta + (1 - p)\hat{P}_\alpha$ and the Poisson point measure $N(p dt, du) + N_\alpha((1 - p) dt, du)$, see (29) and (30). Using these objects we now construct a Feller process X with the help of Itô's procedure. Since $(p\bar{P}^\theta + (1 - p)\hat{P}_\alpha)(1 - e^{-\sigma}) = 1$, the process X has a zero sojourn at 0.

Theorem 4.2. *For each $p \in [0, 1]$, the process X is a (weak) solution to the SDE*

$$X(t) = U_\alpha(t) + S_\theta(pL_0^X(t)), \quad t \geq 0, \quad (37)$$

where U_α is a symmetric α -stable Lévy process, S_θ is a pure-jump Lévy process with the Lévy measure θ , which is independent of U_α , and the process L_0^X is the local time of X at 0.

Remark 4.5. The process U_α appearing in (37) is different from the α -stable Lévy process which is built upon N_α alone.

Remark 4.6. The process $S_\theta(p \cdot)$ has the same distribution as $S_{p\theta}(\cdot)$. As a consequence, X is also a solution to

$$X(t) = U_\alpha(t) + S_{p\theta}(L_0^X(t)), \quad t \geq 0.$$

Theorem C follows from the last remark specialized to $\theta = \eta^*$ (so that $S_\theta = S_\beta$), where θ^* is the measure defined in (7).

Proof of Theorem 4.2. Define the Poisson random measures $N_\eta^{(\varepsilon)}$ and $N^{(\varepsilon)}$ via (33).

Put $\theta^{(\varepsilon)}(dx) := \mathbb{1}_{\{|x|>\varepsilon\}}\theta(dx)$, $\theta_t^{(\varepsilon)} := \theta^{(\varepsilon)}P_t^0$ and

$$q_\varepsilon := \frac{1 - p \int_{|x|>\varepsilon} \mathbb{E}^x(1 - e^{-\sigma})\theta(dx)}{\hat{P}_\alpha(1 - e^{-\sigma})}.$$

Use now Itô's procedure to build the process $X^{(\varepsilon)}$ upon the Poisson point measure $N^{(\varepsilon)}(p dt, du) + N_\alpha(q_\varepsilon dt, du)$, which corresponds to the entrance law $p\theta_t^{(\varepsilon)} + q_\varepsilon\rho_t$ and the characteristic measure $p\bar{P}^{\mathbb{1}_{\{|x|>\varepsilon\}}\theta(dx)} + q_\varepsilon\hat{P}_\alpha$. It follows from the construction that $X^{(\varepsilon)}$ has a zero sojourn at 0, and that

$$X^{(\varepsilon)}(t) = U_\alpha^{(\varepsilon)}(t) + S_\theta^{(\varepsilon)}(p\varphi^{(\varepsilon)}(t)), \quad t \geq 0,$$

where the process $S_\theta^{(\varepsilon)}$ is defined in (34). The remainder of the proof mimics that of Theorem 4.1.

We proceed with a comment.

Assume that we have constructed the process $\tilde{X}^{(\varepsilon)}$ with the help of the Poisson point measure $N^{(\varepsilon)}(p_\varepsilon dt, du) + N_\alpha(q dt, du)$, where

$$p_\varepsilon = \frac{p \int_{\mathbb{R}} \mathbb{E}^x(1 - e^{-\sigma})\theta(dx)}{\int_{|x|>\varepsilon} \mathbb{E}^x(1 - e^{-\sigma})\theta(dx)}.$$

Then

$$\tilde{X}^{(\varepsilon)}(t) = \tilde{U}_\alpha^{(\varepsilon)}(t) + S_\theta^{(\varepsilon)}(p_\varepsilon\tilde{\varphi}^{(\varepsilon)}(t)), \quad t \geq 0.$$

It is likely that the limit relation $\lim_{\varepsilon \rightarrow 0} S_\theta^{(\varepsilon)}(p_\varepsilon\tilde{\varphi}^{(\varepsilon)}(t)) = S_\theta(p\varphi(t))$ a.s. also holds. However, this fact does not follow directly from Lemmas 4.2 and 4.3. Indeed, the composition functional is not continuous on $D \times C_\uparrow([0, \infty))$, where $C_\uparrow([0, \infty))$ is the set of nondecreasing continuous functions, see [37, Section 13.2]. The composition $(S_\theta, \varphi) \mapsto S_\theta \circ \varphi$ would be a.s. continuous, if the function φ were a.s. strictly increasing, which is not the case. The local time φ has many intervals of constancy, which are directly connected to the points of growth of S_θ . \square

4.2. Uniqueness for SDEs with local times for the skew Lévy process.

Theorem 4.3. *Assume that there exists a filtered probability space $(\Omega, (\mathcal{F}_t), \mathcal{F}, \mathbb{P})$, a (\mathcal{F}_t) -adapted symmetric α -stable process U_α with $\alpha \in (1, 2)$ and a homogeneous Feller process $(X, (\mathcal{F}_t))$ with a zero sojourn at 0 which satisfy*

$$X(t) = U_\alpha(t) + H(L_0^X(t)), \quad t \geq 0,$$

where H is a Lévy process of bounded variation, which is independent of U_α . Then there exist $p \in [0, 1]$ and a sigma-finite measure θ on $\mathbb{R} \setminus \{0\}$ satisfying $\mathbb{E}^\theta(1 - e^{-\sigma}) = 1$, for which the processes (X, H) has the same distribution as the processes $(X(\cdot), S_\theta(p \cdot))$ appearing in Theorem 4.2.

Theorem D is a specialization of Theorem 4.3, with the Lévy measure $\theta = \eta$ given by (7) in which case $S_\theta = S_\beta$.

Remark 4.7. Under the assumptions of Theorem 4.3 imposed on (X, U_α) there exists neither a solution to the equation

$$dX(t) = dU_\alpha(t) + \gamma dL_0^X(t), \quad t \geq 0$$

for $\gamma \neq 0$, nor a solution to (37) for $p > 1$.

Proof. Since the local time only grows on the set $\{t \geq 0 : X(t) = 0\}$, the process X is a recurrent extension of a symmetric α -stable Lévy process killed at 0. Hence, there exists a unique value of $p \in [0, 1]$ and a unique measure θ satisfying $E^\theta(1 - e^{-\sigma}) = 1$, for which the entrance law of X is given by $(1 - p)\rho_t + p\theta P_t^0$, see (29) and (30). It follows from Theorem 4.2 that there exists a version \tilde{X} of X satisfying

$$\tilde{X}(t) = \tilde{U}_\alpha(t) + \tilde{S}_\theta(pL_0^{\tilde{X}}(t)), \quad t \geq 0.$$

Observe that the processes X and \tilde{X} are semimartingales. Since the local times $(L_0^X(t))$ and $(L_0^{\tilde{X}}(t))$ are a.s. continuous processes which do not increase a.s. on the sets $\{s \geq 0 : X(s) \neq 0\}$ and $\{s \geq 0 : \tilde{X}(s) \neq 0\}$, respectively, we conclude that, for any $\delta > 0$, with probability 1,

$$\int_{[0, t]} \mathbb{1}_{\{|\tilde{X}(s-)| \geq \delta\}} d(\tilde{S}_\theta(pL_0^{\tilde{X}}(s))) = 0, \quad t > 0$$

and

$$\int_{[0, t]} \mathbb{1}_{\{|X(s-)| \geq \delta\}} d(H(L_0^X(s))) = 0, \quad t \geq 0.$$

This entails

$$\int_{[0, t]} \mathbb{1}_{\{|X(s-)| \geq \delta\}} dX(s) = \int_{[0, t]} \mathbb{1}_{\{|X(s-)| \geq \delta\}} dU_\alpha(s)$$

and

$$\int_{[0, t]} \mathbb{1}_{\{|\tilde{X}(s-)| \geq \delta\}} d\tilde{X}(s) = \int_{[0, t]} \mathbb{1}_{\{|X(s-)| \geq \delta\}} d\tilde{U}_\alpha(s).$$

Sending $\delta \rightarrow 0+$ we conclude that there exists a measurable function $F : D \rightarrow D$ satisfying $\tilde{U}_\alpha = F(\tilde{X})$ and $U_\alpha = F(X)$ a.s. Since the process X has the same distribution as \tilde{X} , there exists a measurable function $G : D \rightarrow D$ satisfying $\tilde{L}_0^{\tilde{X}} = G(\tilde{X})$ and $L_0^X = G(X)$, see [5, Section III.3]. As a consequence, the distributions of the pairs $(H(L_0^{\tilde{X}}), L_0^{\tilde{X}})$ and $(S_\theta(pL_0^X), L_0^X)$ are the same. This, in its turn, ensures that the distributions of the Lévy processes $(H(t))$ and $(S_\theta(pt))$ are the same. \square

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